

Boreal caribou population trend results for Naosap-Reed, Wabowden, The Bog and Charron Lake 2015-2019

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Mark-Recapture Analysis of scat data

- Scat samples analyzed using 9 microsatellite loci
- Caribou-specific Zfx/Zfy primers used for sex id.
- Samples clustered into genotypes and encounter history built
- Encounter histories as formatted for MR analysis by program MARK

```
/* JNP Tonquin 2006-13 K=17 Nhist=156 G=Female Male */  
  
10000100010011000 0 1; /* PCID= 7 */  
10000000000000000 1 0; /* PCID= 10 */  
00000000110000000 0 1; /* PCID= 100 */  
00000101110110011 1 0; /* PCID= 120 */  
  
└──────────┬────────┘  
Encounter history Group counts  
2006 07 08 .... 13 F M
```

Demographic Parameters

Parameter of interest is realised annual growth rate:

$$\lambda_i = N_{i+1} / N_i \text{ and trends in } \lambda$$

Other important demographic parameters are:

N_i = abundance (at survey time in year i)

ϕ_i = (net) survival rate from year i to $i+1$

f_i = (net) fecundity rate

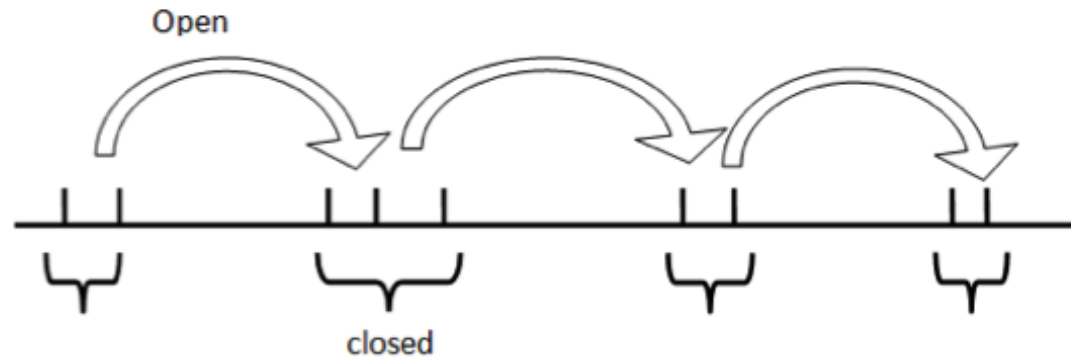
(female calf recruits to N_{i+1} per female in N_i)

Parameter estimates for each sex.

Note that $\lambda_i = \phi_i + f_i$

Robust Design Models (RD)

- Annual (primary) sampling occasions involving multiple within-year surveys (secondary samples)



- Population assumed “open” among years (subject to recruitment and losses)
- Population assumed closed across secondary samples and samples independent
 - Note trade-off: if 2^{ndary} samples too close in time, sighting independence fails; if too far apart, closure fails

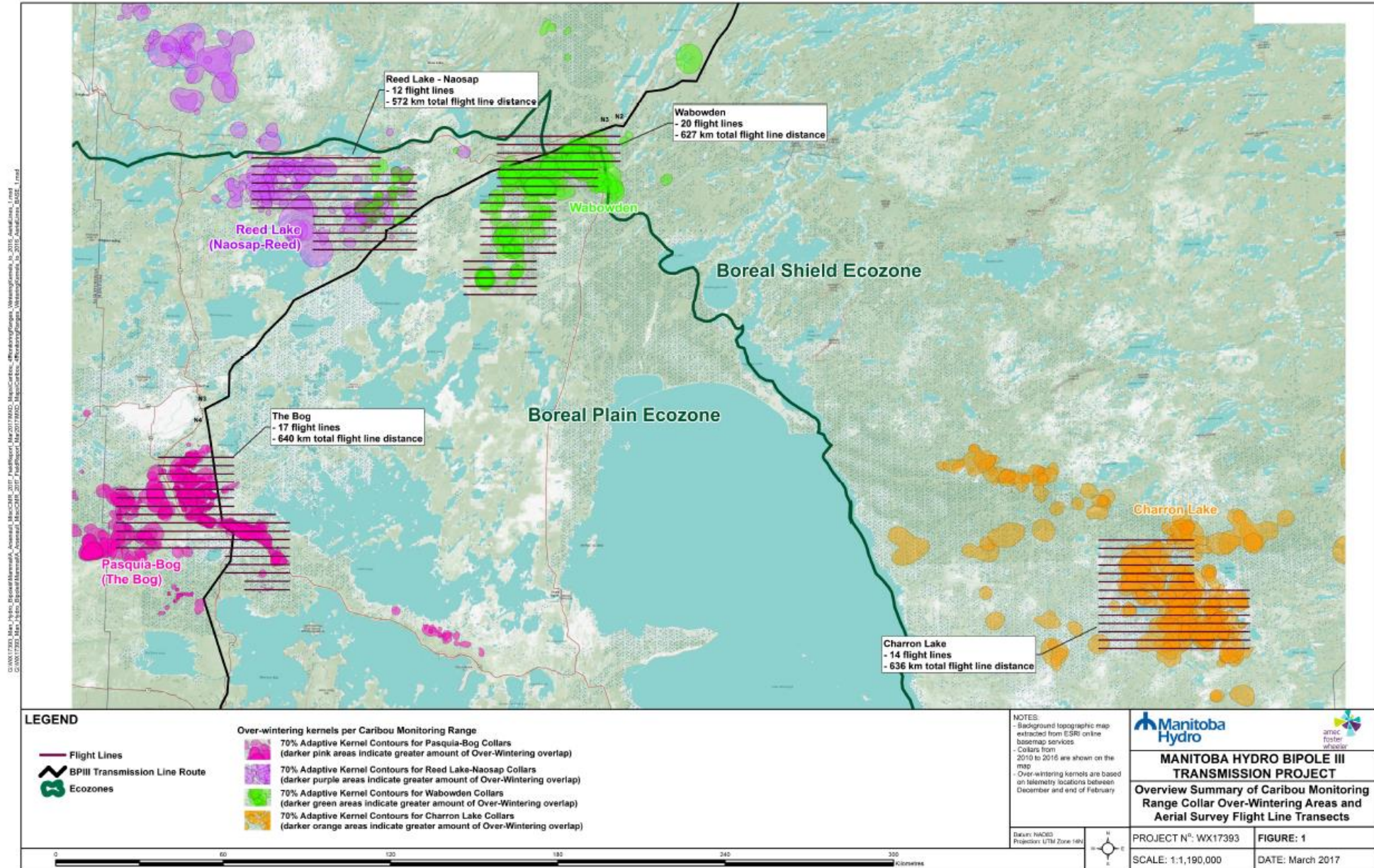
RD parameter estimation

- Requires a minimum of 2 secondary samples per primary
- Pradel models (Pradel 1996, Cooch and White 2015) applied to primary intervals:
 - Provides estimates of λ_i , ϕ_i , f_i
 - All rates are normalized to annual rates by providing deltas ($\delta_i = t_{i+1} - t_i$) in years; Allows for unequal δ_i
- Closed models applied to secondary samples
 - Provides estimates of N_i
 - If only 2 secondaries per primary, model choices are (Otis et al. 1978)
 M_0 , M_t , M_b (“behavioural” model: marked status affects capture rate)
- Capture rate (nuisance parameter p_i) estimated from all data
 - Importance of reducing number of p_i by using constraints and/or covariates.
- CJS models (ignoring RD structure) can be used to model p and ϕ and are used to test closure, judge constraints on p and ϕ

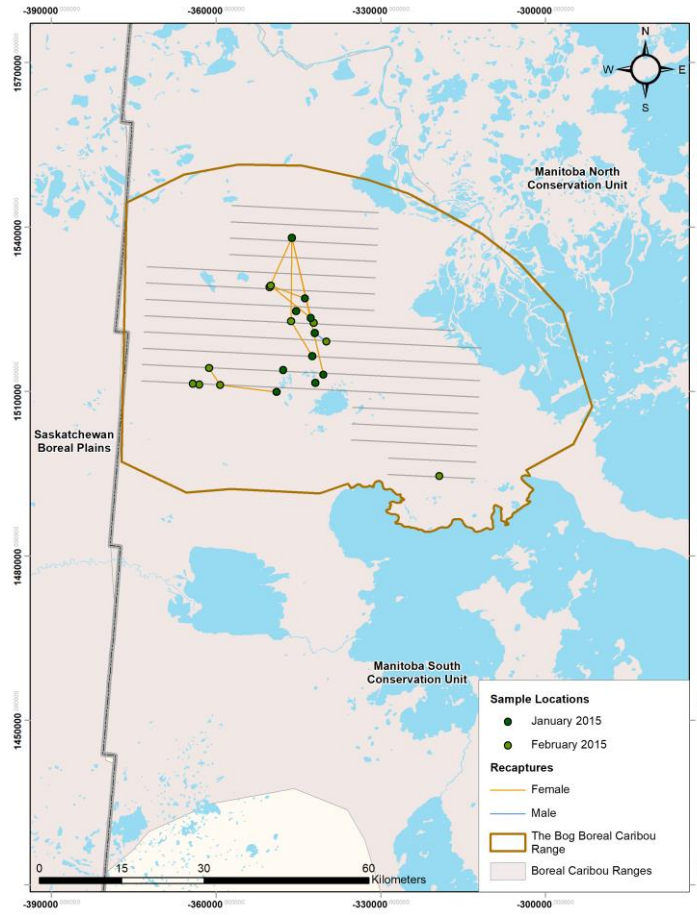
RD analyses in MARK

- Fit and rank systematic battery of models: Pradel primary + Closed (Otis) secondary models.
 - $\phi: (g \times t), (g), (t), (\cdot)$
 - $\lambda: (g \times t), (g), (t), (\cdot)$
 - $p: (g \times t), (t), (g \times \text{effort})$
 - $c: (g \times t), (t), \{c=p_2\}$ (no within-year capture effect)
 - $F_0: (g \times t)$ (number of never-sited animals; note that N is a derived parameter)
- Obtain model averaged estimates for N , λ , ϕ . Top models weighted by AICc (a score combining GOF with number of model parameters).
- Fit constrained models with $\lambda(g)$, $\lambda(t)$, *and* $\lambda(\cdot)$ to estimate sex, time, and overall rate of change (these models may not be among highest ranked models but provide best estimates of average λ over these subsets).

Manitoba Hydro CMR scat collection 2015-2019

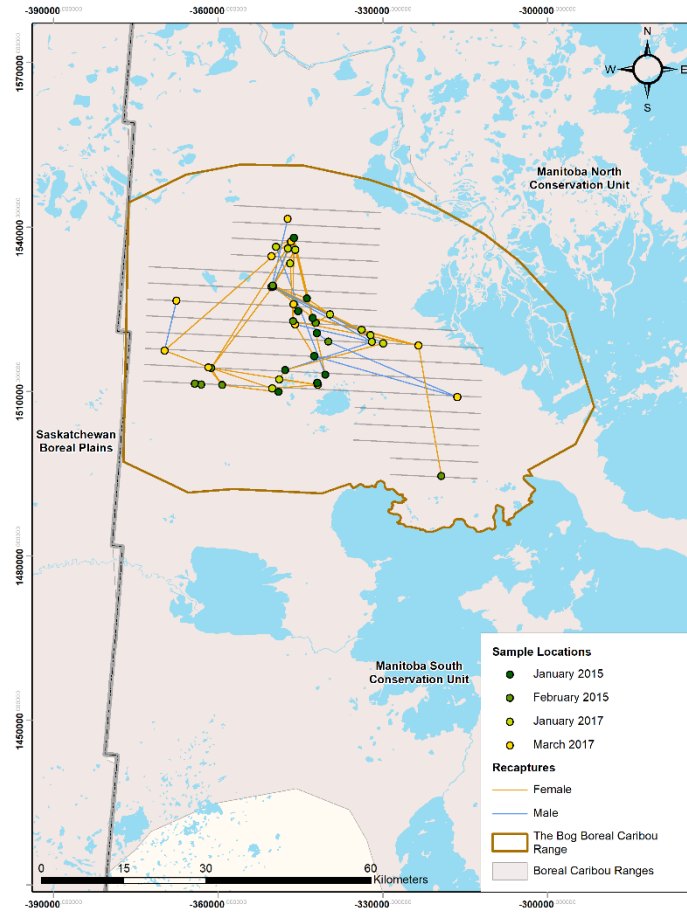


The Bog Scat Collection – 2015, 2017 & 2019



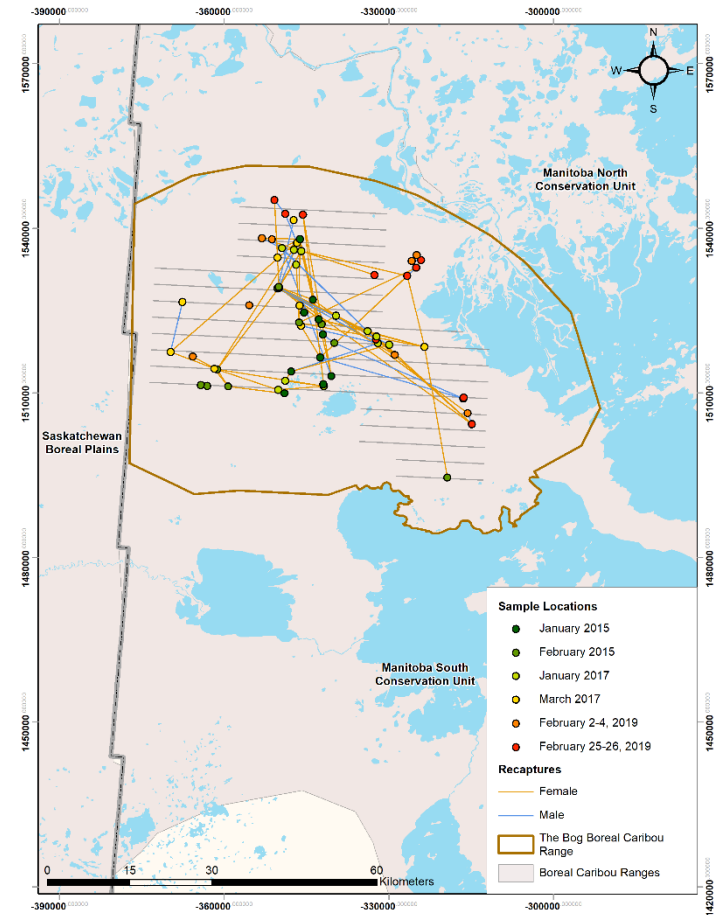
- 183 samples successfully profiled
 - Occasion 1 – 81 samples, Occasion 2 – 102 samples
- 88 unique genotypes – 72 females, 16 males
 - Occasion 1 – 47 genotypes, Occasion 2 – 41 genotypes
- 16/41 (39%) genotypes recaptured between January and February

2015



- 243 samples successfully profiled
 - Occasion 3 – 110 samples, Occasion 4 – 133 samples
- 97 unique genotypes – 62 females, 35 males
 - Occasion 3 – 56 genotypes, Occasion 4 – 41 genotypes
- 10/41 (24%) genotypes recaptured between January and March

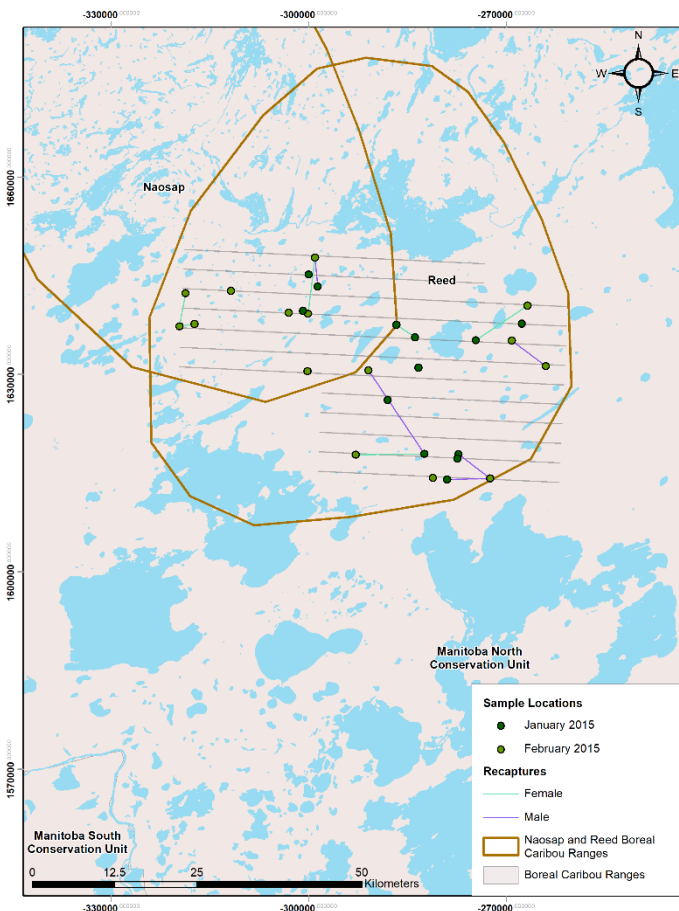
2017



- 203 samples successfully profiled
 - Occasion 5 – 95 samples, Occasion 6 – 108 samples
- 60 unique genotypes – 40 females, 20 males
 - Occasion 5 – 25 genotypes, Occasion 6 – 35 genotypes
- 11/35 (31%) genotypes recaptured between early and late February

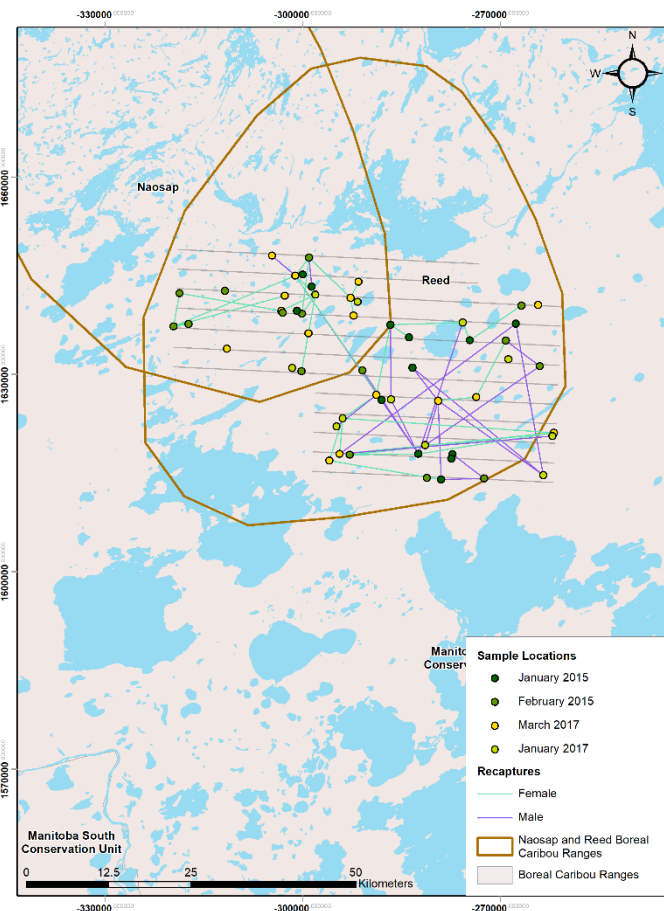
2019

The Naosap-Reed Scat Collection – 2015, 2017 & 2019



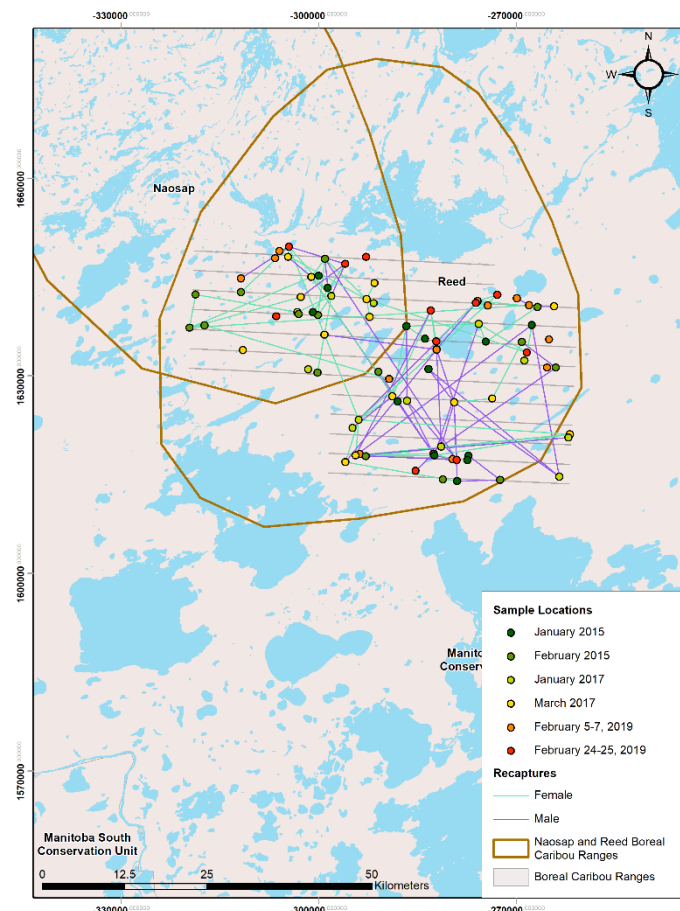
- 225 samples successfully profiled
 - Occasion 1 – 91 samples, Occasion 2 – 134 samples
- 109 unique genotypes – 57 females, 52 males
 - Occasion 1 – 49 genotypes, Occasion 2 – 60 genotypes
- 10/60 (17%) genotypes recaptured between January and February

2015



- 298 samples successfully profiled
 - Occasion 3 – 86 samples, Occasion 4 – 212 samples
- 143 unique genotypes – 82 females, 61 males
 - Occasion 3 – 50 genotypes, Occasion 4 – 93 genotypes
- 13/93 (14%) genotypes recaptured between January and March

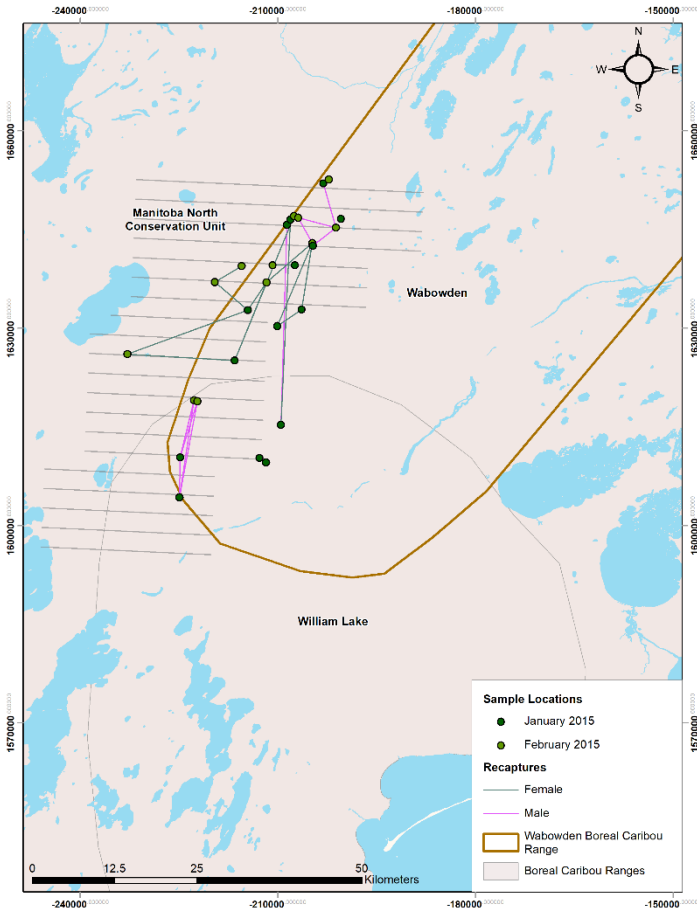
2017



- 225 samples successfully profiled
 - Occasion 5 – 91 samples, Occasion 6 – 134 samples
- 118 unique genotypes – 75 females, 43 males
 - Occasion 5 – 58 genotypes, Occasion 6 – 60 genotypes
- 20/60 (33%) genotypes recaptured between early and late February

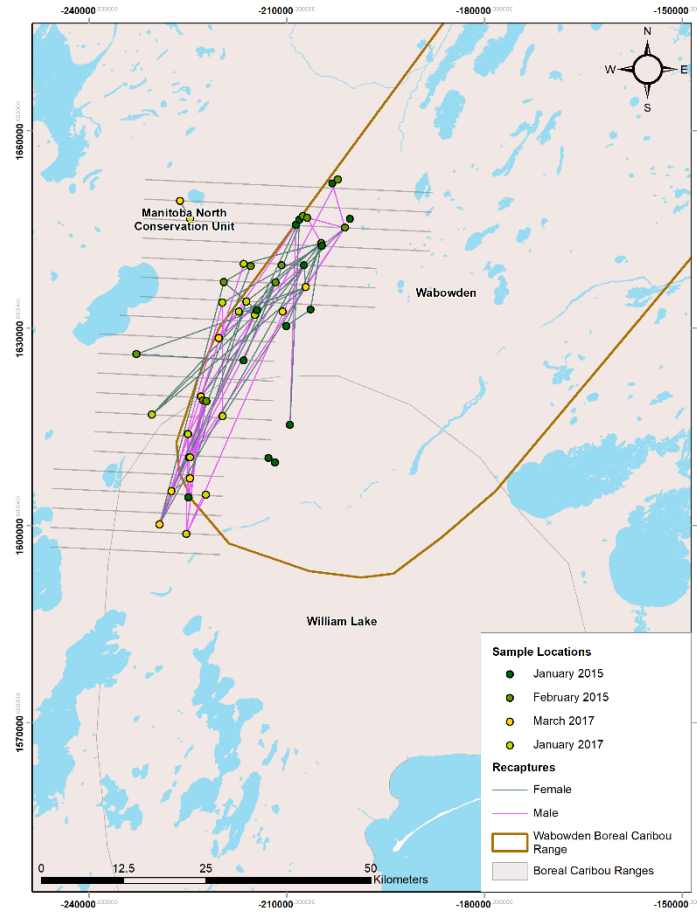
2019

Wabowden Scat Collection – 2015, 2017 & 2019



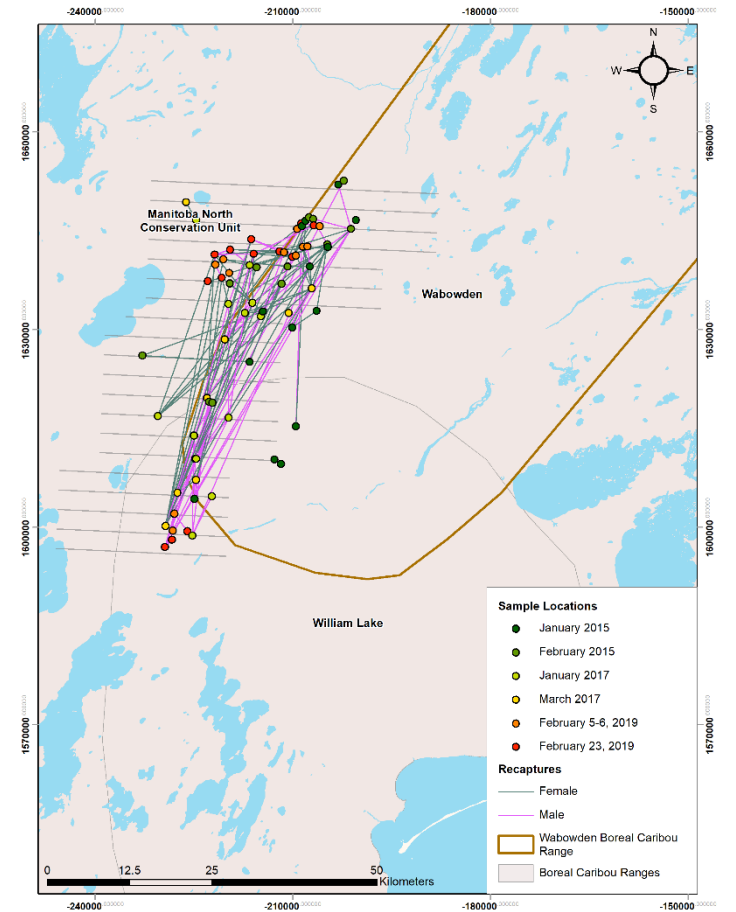
- 208 samples successfully profiled
 - Occasion 1 – 120 samples, Occasion 2 – 88 samples
- 108 unique genotypes – 64 females, 44 males
 - Occasion 1 – 64 genotypes, Occasion 2 – 44 genotypes
- 26/44 (59%) genotypes recaptured between January and February

2015



- 208 samples successfully profiled
 - Occasion 3 – 83 samples, Occasion 4 – 125 samples
- 98 unique genotypes – 55 females, 43 males
 - Occasion 3 – 47 genotypes, Occasion 4 – 51 genotypes
- 15/51 (29%) genotypes recaptured between January and March

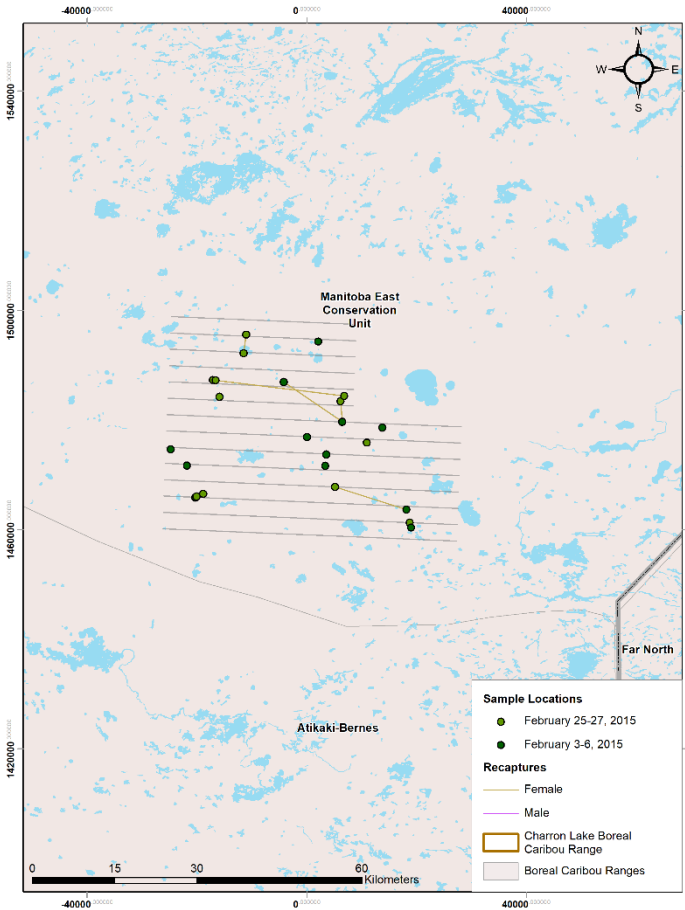
2017



- 230 samples successfully profiled
 - Occasion 5 – 102 samples, Occasion 6 – 128 samples
- 97 unique genotypes – 52 females, 45 males
 - Occasion 5 – 39 genotypes, Occasion 6 – 58 genotypes
- 18/58 (31%) genotypes recaptured between early and late February

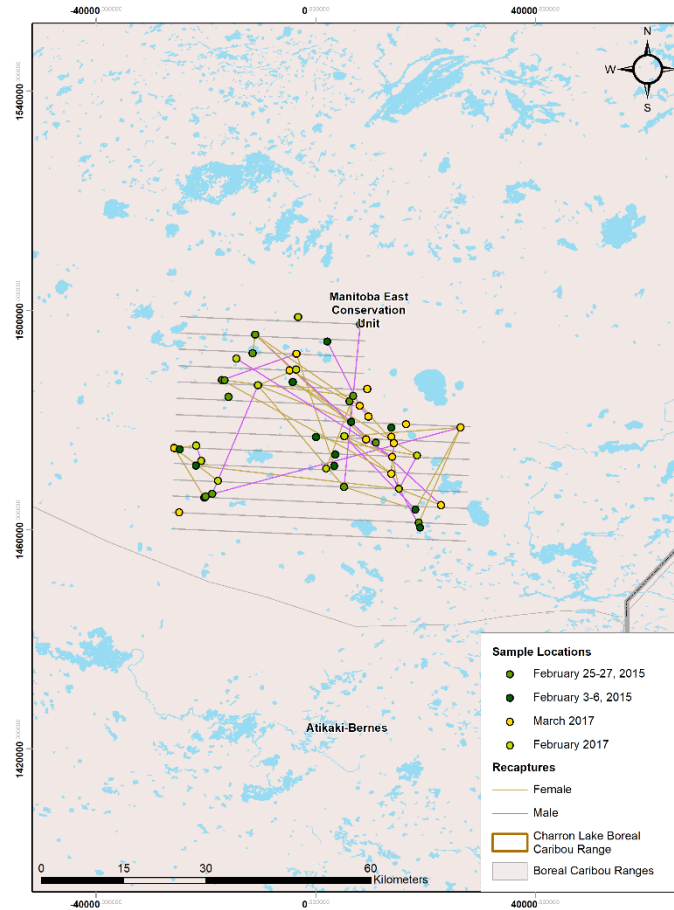
2019

Charron Lake Scat Collection – 2015, 2017 & 2019



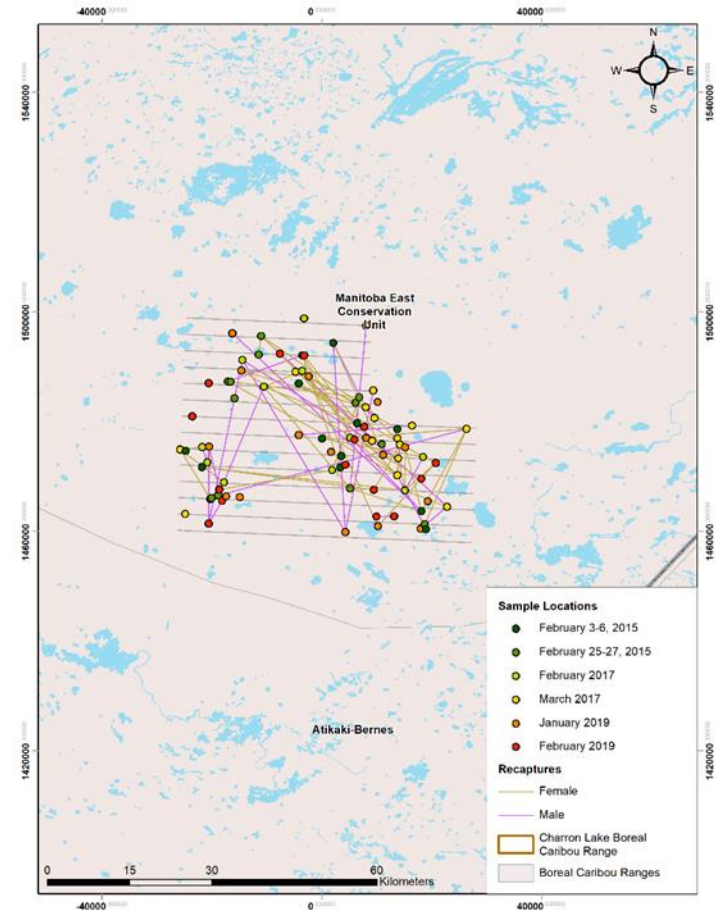
- 225 samples successfully profiled
 - Occasion 1 – 126 samples, Occasion 2 – 99 samples
- 131 unique genotypes – 98 females, 33 males
 - Occasion 1 – 77 genotypes, Occasion 2 – 54 genotypes
- 5 (9%) genotypes recaptured between early and late February

2015



- 279 samples successfully profiled
 - Occasion 3 – 147 samples, Occasion 4 – 132 samples
- 177 unique genotypes – 163 females, 64 males
 - Occasion 3 – 91 genotypes, Occasion 4 – 86 genotypes
- 9/86 (10%) genotypes recaptured between February and March

2017



- 271 samples successfully profiled
 - Occasion 5 – 125 samples, Occasion 6 – 146 samples
- 170 unique genotypes – 111 females, 59 males
 - Occasion 5 – 76 genotypes, Occasion 6 – 94 genotypes
- 6/94 (6%) genotypes recaptured between January and February

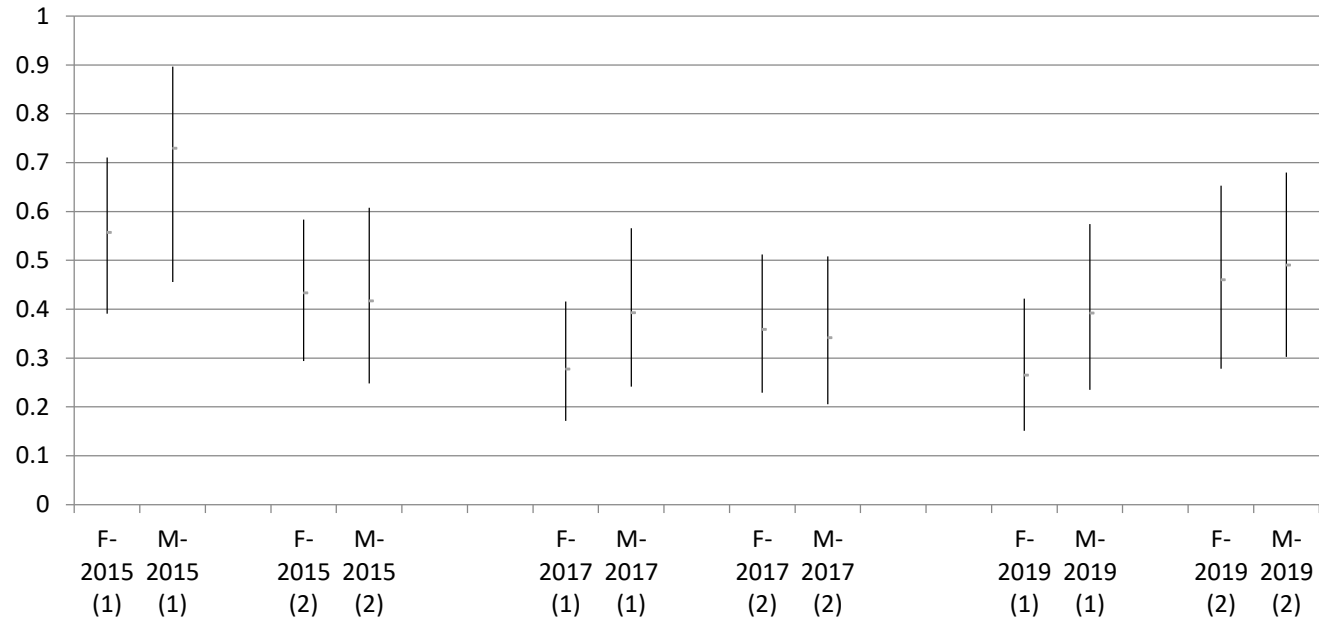
2019

Model Fits Wabowden 2015-2019

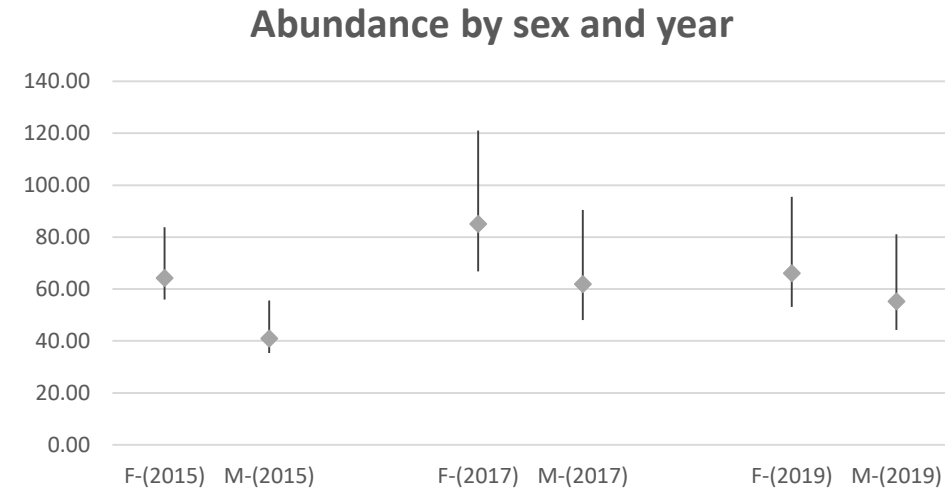
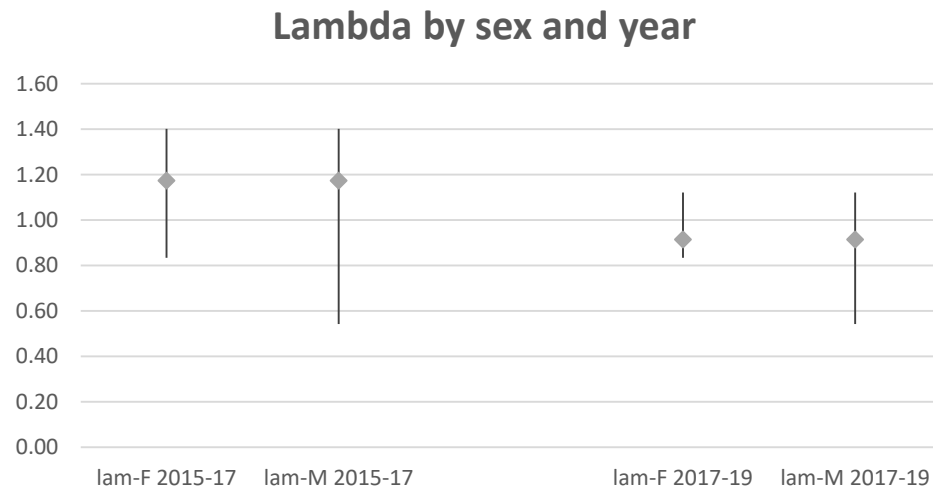
- Considerable variation in effort over sample times and $\text{logit}(p)$ was well explained by a linear model in effort.
- With a more general p model, it appeared that both Φ and Lam also show time trends. Both Φ and Lam show a decrease from 2015-17 to 2017-19 and this is consistent with a constant recruitment as most of the decline in Lam is explained by the decline in survival Φ
- The declines are not big enough to be statistically significant given the estimation uncertainty and both Lam confidence intervals overlap 1.0.
- In fact, the top model (#1 on Models tab) supports no time effects on Φ or Lambda , but there is no biological reason why this restriction should apply and the #1 model is too restrictive. To avoid bias, the more general model (#4) $\{\Phi(t) \text{ Lam}(t) p(t) p=c\}$ was adopted as the best model for forming estimates.

#	Model	AICc	Delta AICc	AICc Weights	Model Likelihood	Num. Par	Deviance
1	$\{\Phi(.) \text{ Lam}(.) p(t) p=c \text{ PIM}\}$	-152.603	0	0.32413	1	14	-1241.01
2	$\{\Phi(.) \text{ Lam}(t) p(t) p=c \text{ PIM}\}$	-151.606	0.9977	0.19682	0.6072	15	-1242.23
3	$\{\Phi(t) \text{ Lambda}(t) p(a + b\text{EFF}) c=p f_0(g^*t) \text{ DM}\}$	-150.831	1.7722	0.13363	0.4123	12	-1234.86
4	$\{\Phi(t) \text{ Lam}(t) p(t) p=c \text{ PIM}\}$	-150.475	2.1288	0.11181	0.3449	16	-1243.33
5	$\{\Phi(t) \text{ Lambda}(t) p(t) c=p f_0(g^*t) \text{ DM}\}$	-150.475	2.1288	0.11181	0.3449	16	-1243.33
6	$\{\Phi(t) \text{ Lam}(.) p(t) p=c \text{ PIM}\}$	-150.394	2.2093	0.10739	0.3313	15	-1241.02
7	$\{\Phi(.) \text{ Lam}(.) p(g) p=c\}$	-145.202	7.4009	0.00801	0.0247	10	-1224.91
8	$\{\Phi(g^*t) \text{ Lam}(g^*t) p(t) p=c \text{ PIM}\}$	-142.335	10.2681	0.00191	0.0059	20	-1244.26
9	$\{\Phi(t) \text{ Lam}(t) p(g^*t) p=c \text{ PIM}\}$	-141.697	10.9068	0.00139	0.0043	22	-1248.26
10	$\{\Phi(g) \text{ Lam}(g) p(g) p=c\}$	-141.648	10.9555	0.00135	0.0042	12	-1225.67
11	$\{\Phi(t) \text{ Lam}(t) p(g) p=c\}$	-141.431	11.1726	0.00122	0.0038	12	-1225.46
12	$\{\Phi(t) \text{ Lam}(g^*t) p(g^*t) p=c \text{ PIM}\}$	-138.847	13.7567	0.00033	0.001	24	-1250.11
13	$\{\Phi(g^*t) \text{ Lam}(t) p(g^*t) p=c \text{ PIM}\}$	-137.438	15.1658	0.00017	0.0005	24	-1248.7
14	$\{\Phi(g^*t) \text{ Lam}(g^*t) p(g^*t) p=c \text{ PIM}\}$	-134.11	18.4932	0.00003	0.0001	26	-1250.15
15	$\{\Phi(t) \text{ Lam}(t) p(t) c(t) \text{ PIM fail}\}$	40.6855	193.2888	0	0	19	0
16	$\{\Phi(t) \text{ Lambda}(t) p(t) c(t) f_0(g^*t) \text{ DM fail}\}$	40.6855	193.2888	0	0	19	0

Results Wabowden: Capture rate



Results Wabowden: Abundance and lambda



For both sexes, populations rose between 2015 and 17 and declined between 2017 and 19. This is also reflected in the lambda values which are above 1 for 2015-17 and below 1 for 2017-19. All estimates overlapped 1.0. Survival rates were 0.77 (0.06) in 2015-17 and 0.66 (0.07) in 2017-2019. Pregnancy rates (from hormones analysis) averaged 0.78 (0.86 (2015), 0.80 (2017), 0.70 (2019)).

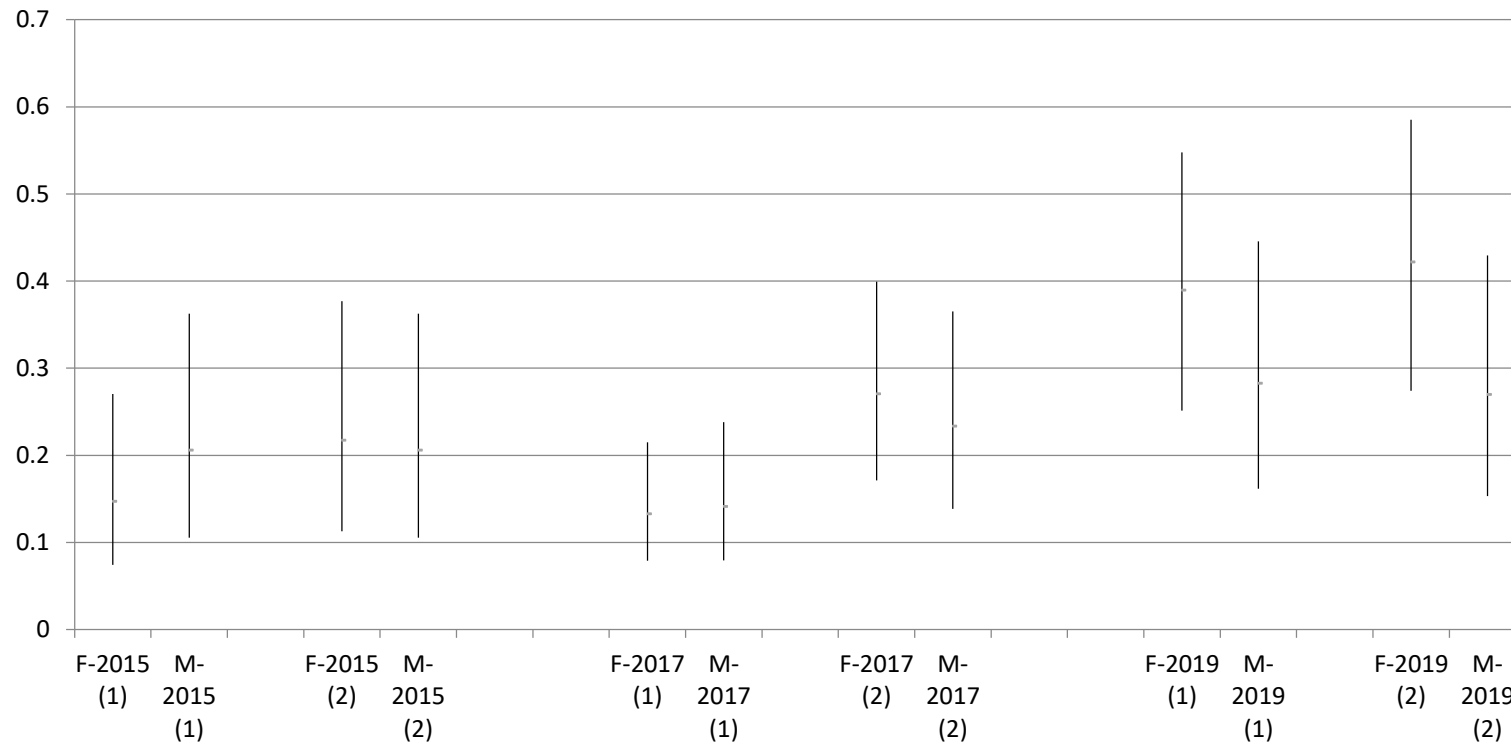
With the more general best model the precision of the N estimates from RD analysis are much the same as from the Peterson estimates within years and sexes.

Model Fits Naosap-Reed 2015-2019

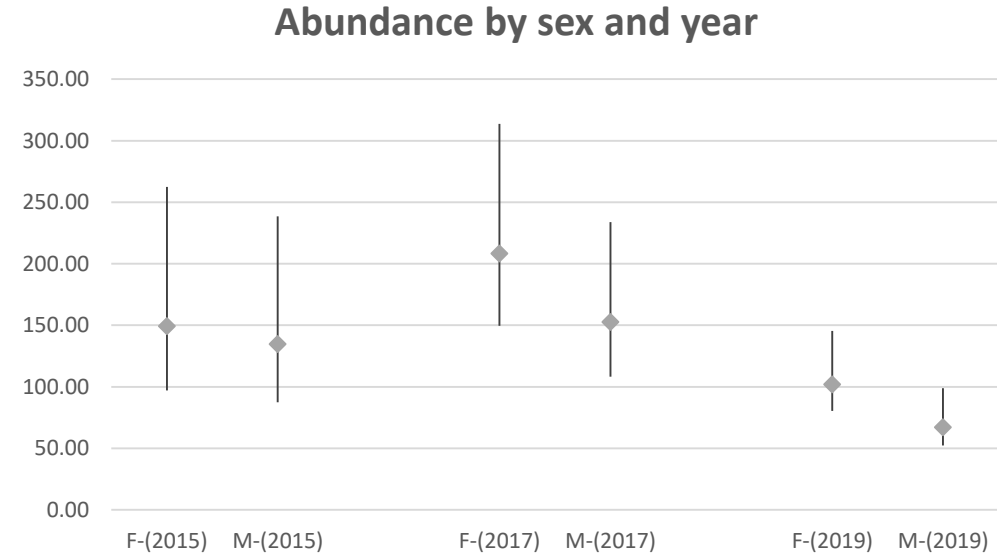
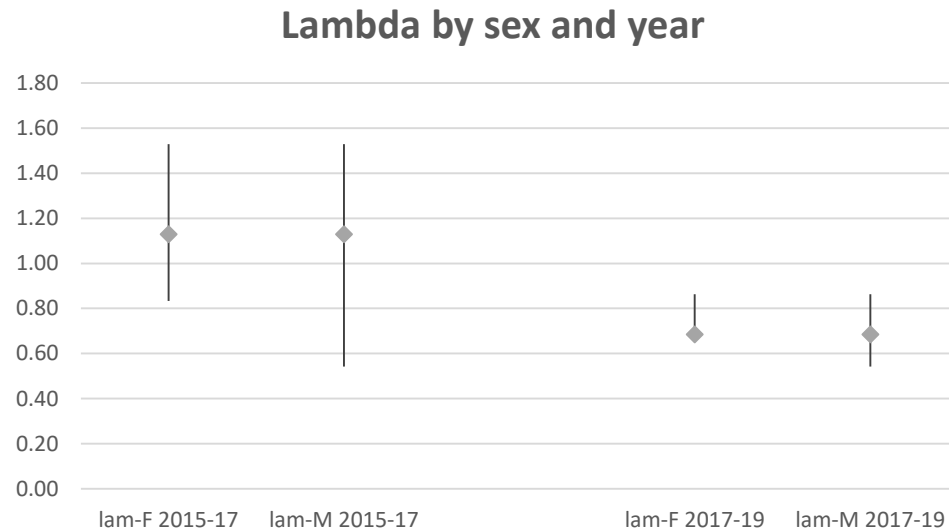
- Model conclusions are similar to Wabowden although this population is doing much more poorly. CJS analyses support a time effect on p but no group (sex) effect on Φ or p .
- RD analyses support a constant $\Phi(.)$ and/or $\text{Lam}(.)$ but opted for $\Phi(t)$ $\text{Lam}(t)$ $p(t)$ as the best model to get a better idea of time trends. The more restrictive $\Phi(.)$ and/or $\text{Lam}(.)$ models are supported only because there isn't sufficient precision to rule them out.

Model	AICc	Delta AICc	AICc Weights	Model Likelihood	Num. Par	Deviance
{ $\Phi(.)$ $\text{Lam}(.)$ $p(t)$ $p=c$ PIM}	-456.958	0	0.91309	1	8	-1873.33
{ $\Phi(.)$ $\text{Lam}(.)$ $p(g)$ $p=c$ }	-451.239	5.7193	0.05231	0.0573	4	-1859.32
{ $\Phi(t)$ $\text{Lam}(t)$ $p(t)$ $p=c$ PIM}	-448.959	7.9991	0.01673	0.0183	16	-1882.47
{ $\Phi(t)$ $\text{Lam}(.)$ $p(t)$ $p=c$ PIM}	-448.7	8.2586	0.0147	0.0161	14	-1877.85
{ $\Phi(t)$ $\text{Lam}(t)$ $p(g)$ $p=c$ }	-444.359	12.5993	0.00168	0.0018	12	-1869.2
{ $\Phi(.)$ $\text{Lam}(t)$ $p(t)$ $p=c$ PIM}	-442.438	14.5207	0.00064	0.0007	15	-1873.76
{ $\Phi(t)$ $\text{Lam}(t)$ $p(g^*t)$ $p=c$ PIM}	-441.273	15.6852	0.00036	0.0004	22	-1888.16
{ $\Phi(g^*t)$ $\text{Lam}(g^*t)$ $p(g^*t)$ $p=c$ PIM}	-440.206	16.7519	0.00021	0.0002	25	-1893.96
{ $\Phi(g^*t)$ $\text{Lam}(t)$ $p(g^*t)$ $p=c$ PIM}	-440.167	16.7913	0.00021	0.0002	24	-1891.62
{ $\Phi(t)$ $\text{Lam}(g^*t)$ $p(g^*t)$ $p=c$ PIM}	-436.912	20.0467	0.00004	0	24	-1888.36
{ $\Phi(g)$ $\text{Lam}(g)$ $p(g)$ $p=c$ }	-436.602	20.3568	0.00003	0	12	-1861.45
{ $\Phi(t)$ $\text{Lam}(t)$ $p(t)$ $c(t)$ PIM fail}	40.1714	497.1297	0	0	19	0

Results Naosap-Reed: Capture rate



Results Naosap-Reed: Abundance and λ



Estimates for N change substantially in the $\Phi(.)$ and/or $\lambda(.)$ models so it is better to go with the more general model as it is likely to be less biased.

CV's for N estimates mostly around 30% for N-Reed whereas they were around 15% for Wabowden. Same trend as with Wabowden: numbers for both sexes seem to increase between 2015 and 2017 and then decline between 2017 and 2019.

$\lambda > 1$ in first interval and $\lambda < 1$ in second. Big decline in survival in second interval leading to big drop in λ : $\Phi(2015-17) = 0.89 (0.08)$, $\Phi(2017-19) = 0.55 (0.06)$. $\lambda(2017-19)$ is significantly below 1. Large drop in λ is not entirely accounted for by decline in Φ so there must have been recruitment failure too. Pregnancy rates (from hormones analysis) averaged 0.72 over the 3 years; being lower in 2017 (0.79 (2015), 0.65 (2017), 0.71 (2019)).

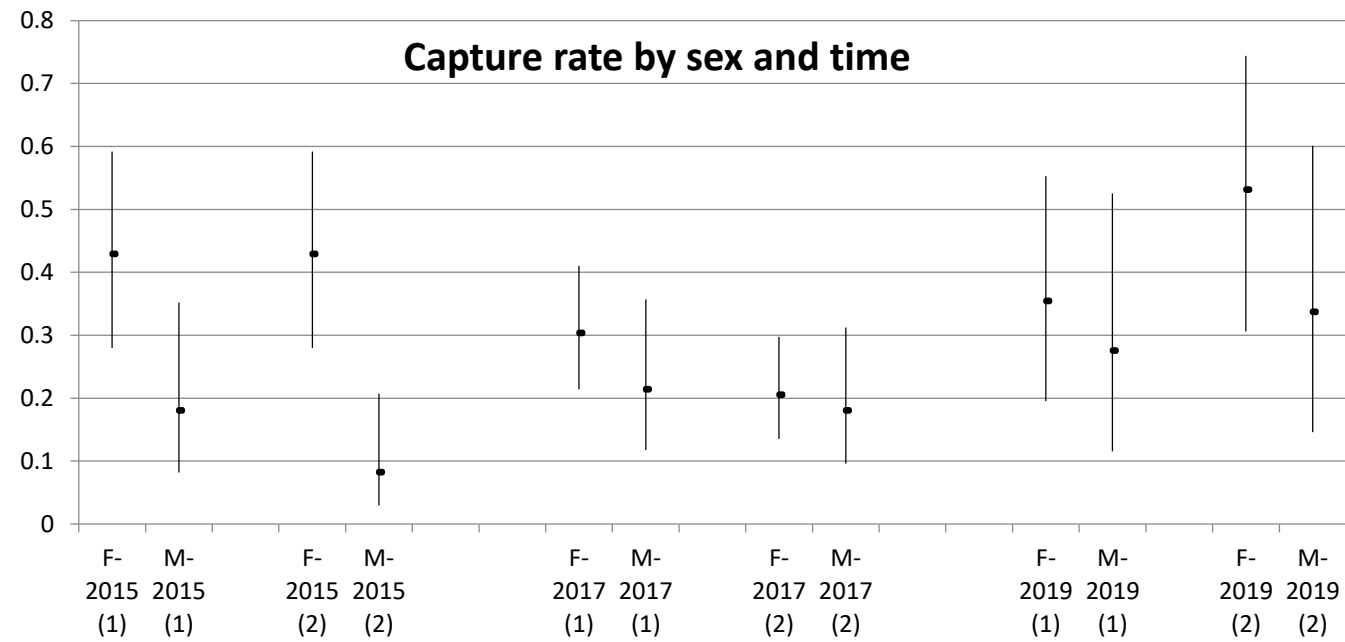
Model Fits The Bog 2015-2019

- The models used were the same as for Naosap-Reed and results are generally similar though specifically less “clean” and precise.
- The CJS models showed a group and time effect on p and a time effect but no group effect on Φ (the top 4 models were all $\Phi(t)$; models with $\Phi(g)$ or $\Phi(g*t)$ had AIC weights below 5%). The group effect on p is almost solely due to very low capture rates of males relative to females in 2015.
- Males show a consistently lower capture rate relative to females in all years.

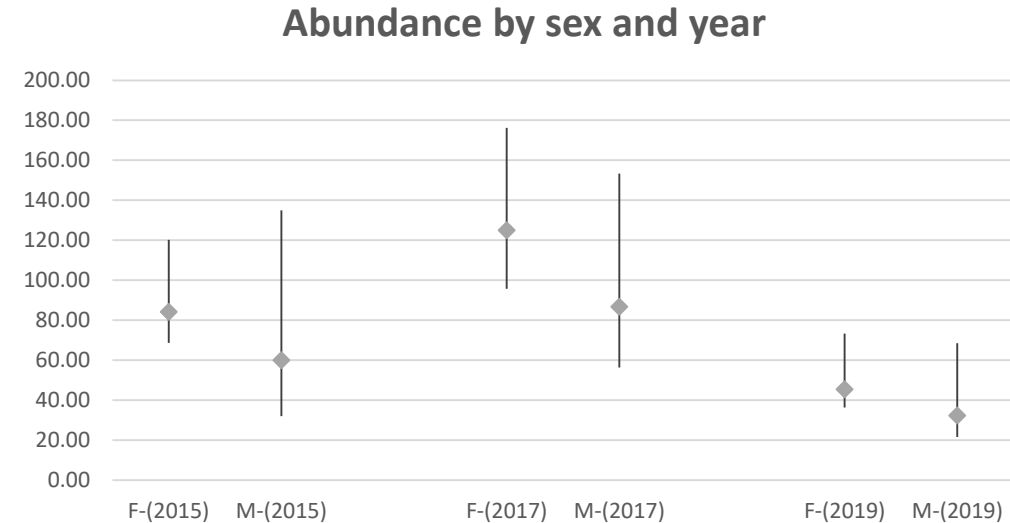
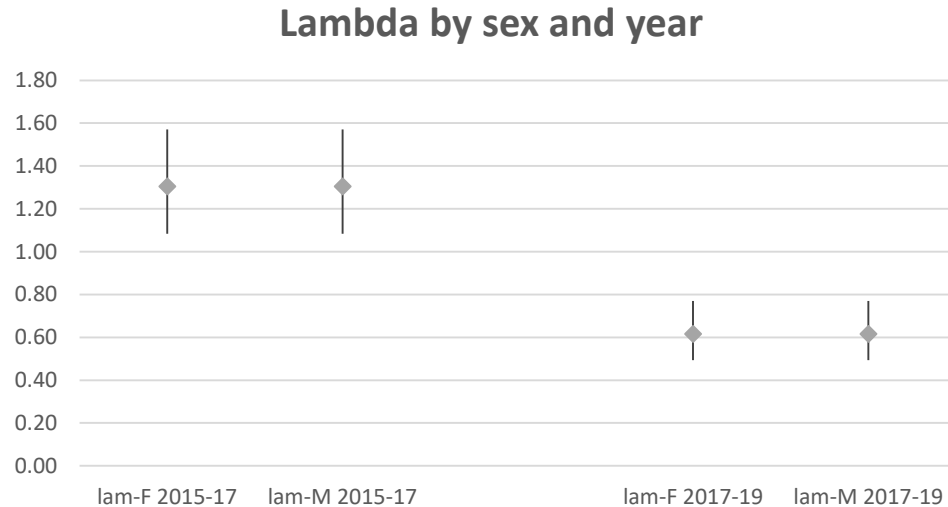
#	Model	AICc	Delta AICc	AICc Weights	Model Likelihood	Num. Par	Deviance	-2log(L)
1	{ $\Phi(t)$ $\text{Lam}(t)$ $p(g*t)$ $p=c$ PIM}	-131.28	0.00	0.61	1.00	21.00	-1019.15	-177.43
2	{ $\Phi(g*t)$ $\text{Lam}(t)$ $p(g*t)$ $p=c$ PIM}	-128.87	2.42	0.18	0.30	23.00	-1021.58	-179.87
3	{ $\Phi(t)$ $\text{Lam}(g*t)$ $p(g*t)$ $p=c$ PIM}	-127.76	3.53	0.10	0.17	23.00	-1020.47	-178.75
4	{ $\Phi(t)$ $\text{Lam}(t)$ $p(t)$ $p=c$ PIM}	-126.87	4.41	0.07	0.11	15.00	-1000.69	-158.97
5	{ $\Phi(g*t)$ $\text{Lam}(g*t)$ $p(g*t)$ $c=p$ F0($g*t$) PIM}	-124.17	7.12	0.02	0.03	25.00	-1021.82	-180.10
6	{ $\Phi(t)$ $\text{Lam}(t)$ $p(g)$ $p=c$ }	-123.39	7.89	0.01	0.02	12.00	-990.46	-148.74
7	{ $\Phi(g)$ $\text{Lam}(g)$ $p(g)$ $p=c$ }	-121.17	10.11	0.00	0.01	12.00	-988.23	-146.52
8	{ $\Phi(.)$ $\text{Lam}(.)$ $p(t)$ $p=c$ PIM}	-114.41	16.87	0.00	0.00	14.00	-985.96	-144.24
9	{ $\Phi(.)$ $\text{Lam}(t)$ $p(t)$ $p=c$ PIM}	-113.88	17.41	0.00	0.00	15.00	-987.69	-145.97
10	{ $\Phi(t)$ $\text{Lam}(.)$ $p(t)$ $p=c$ PIM}	-113.53	17.75	0.00	0.00	15.00	-987.35	-145.63
11	{ $\Phi(.)$ $\text{Lam}(.)$ $p(g)$ $p=c$ }	-110.61	20.67	0.00	0.00	10.00	-973.27	-131.55

- All top models in both the CJS models and RD models support a time effect on Φ but no sex effect. However, all models gave an estimate of $\Phi(2015 \text{ to } 2017)$ of 1.0 and this parameter had to be fixed at 1.0 to avoid convergence problems. Similarly, there was support for a strong time effect on Lam but no sex effect. So the top model { $\Phi(t)$ $\text{Lam}(t)$ $p(g*t)$ $p=c$ } was chosen for all the estimates.

Results The Bog: Capture rate



Results The Bog: Abundance and λ



As with N-Reed, Lambda is high over 2015-17 and then crashed to less than half the value over 2017-19.

There is a large decrease in Phi as well: $\Phi(2015-17) = 1.0$; $\Phi(2017-19) = 0.54$ (0.07), but again as with N-Reed, not enough to fully explain the decline in Lambda, so some recruitment failure is implied. Pregnancy rates (from hormones analysis) averaged 0.74; being higher in 2017 (0.68 (2015), 0.80 (2017), 0.72 (2019)).

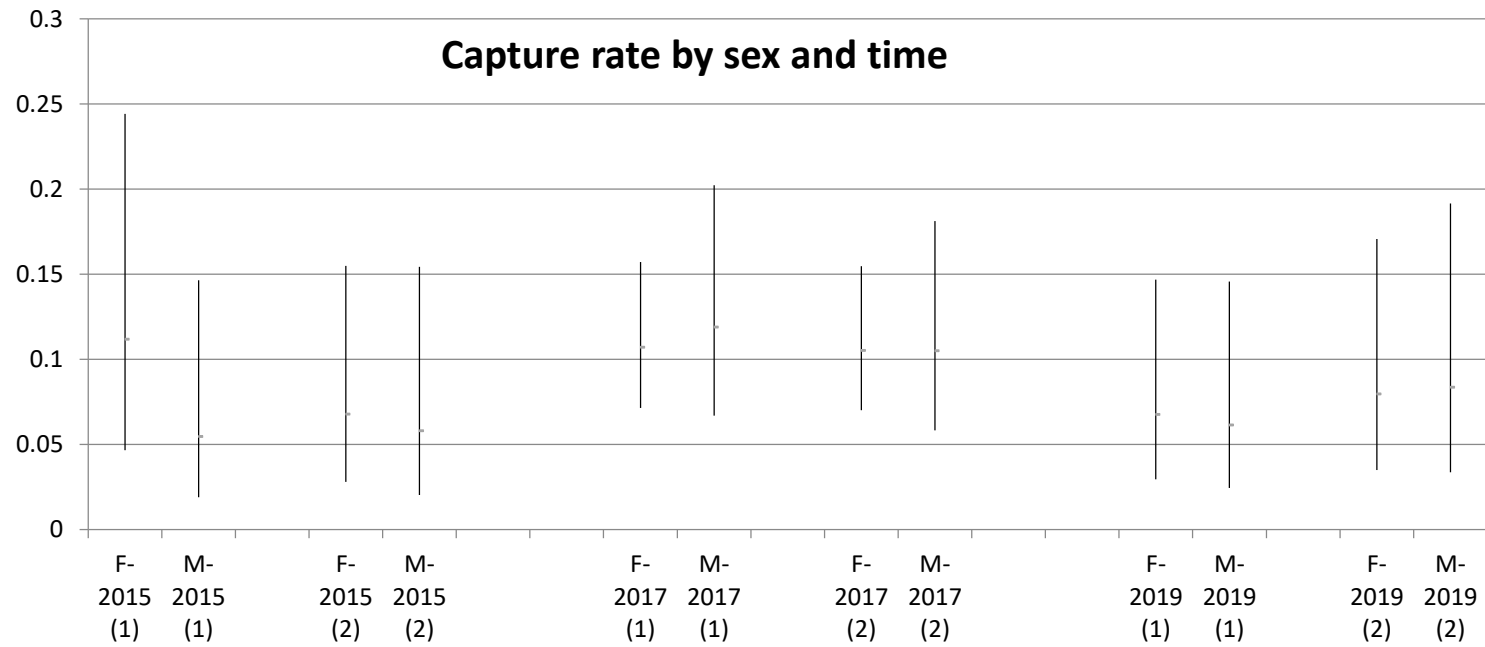
Lambda Confidence intervals do not overlap 1.0 in either time interval so $\Lambda > 1$ in 2015-17 and $\Lambda < 1$ in 2017-19 are both significantly different from stable rates of change ($\Lambda = 1$).

Model Fits Charron Lake 2015-2019

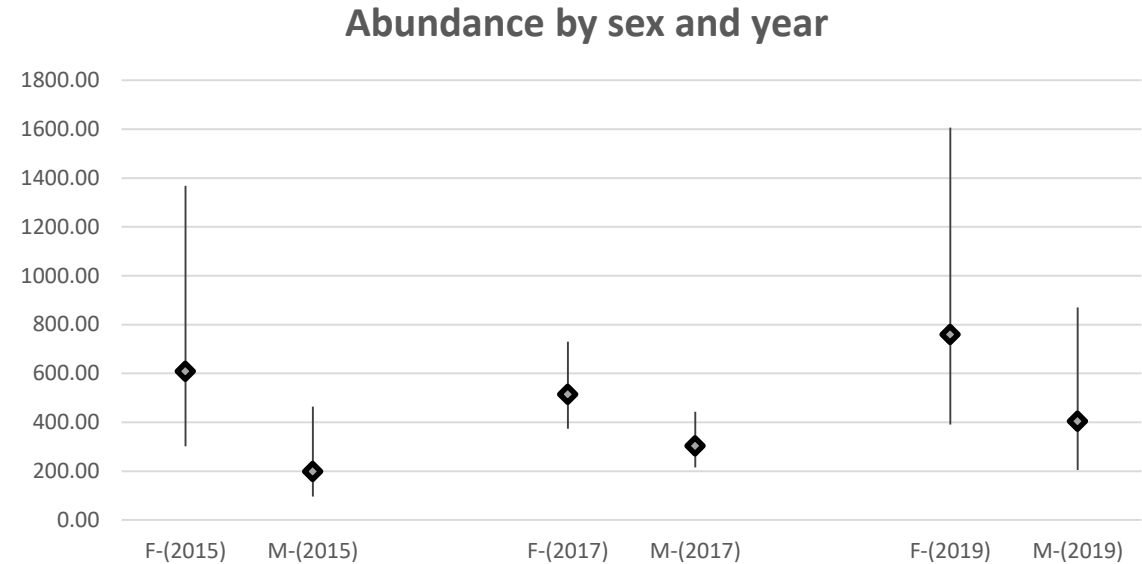
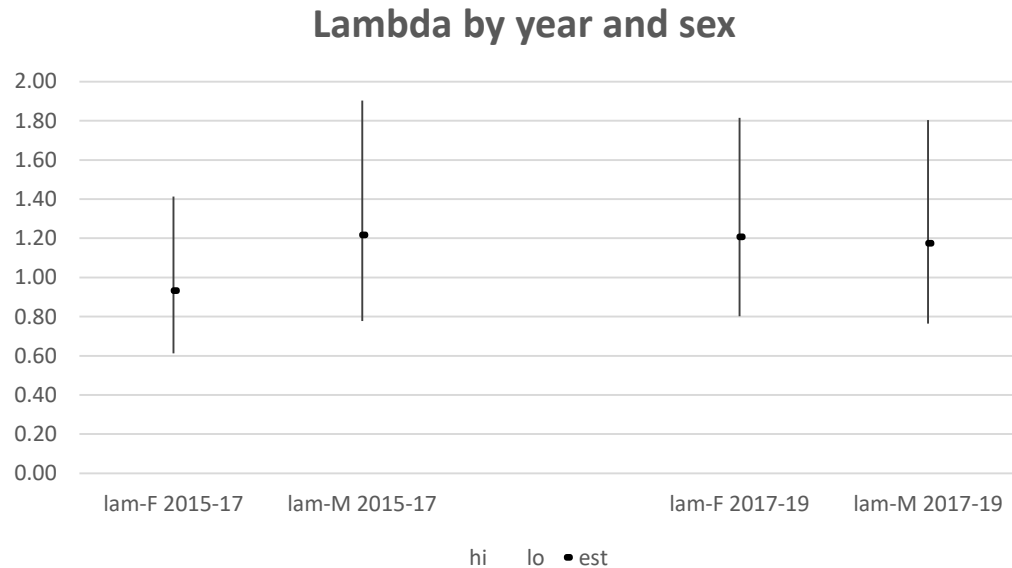
- Models used were as for the Bog with some additional models to explore time and sex effects on Φ and λ . Both the CJS and RD models support a time effect on p but no sex effect and the highest ranking model is $\Phi(t) \lambda(t) p(t)$.
- There is visual and model weight support for no time or sex effect on λ . There was very little variation in sampling effort over sample times and so effort was not a good explanatory covariate for p .

#	Model	AICc	Delta AICc	AICc Weights	Model Likelihood d	Num. Par	Deviance	-2ln(L)
1	{ $\Phi(t) \lambda(t) p(t) p=c$ PIM}	-1181.8	0.00	0.67	1.00	9	-3031.7	-1200.2
2	{ $\Phi(t) \lambda(.) p(t) p=c$ PIM}	-1179.5	2.39	0.20	0.30	10	-3031.4	-1199.9
3	{ $\Phi(g^*t) \lambda(g^*t) p(t) p=c$ PIM}	-1176.0	5.83	0.04	0.05	15	-3038.5	-1207.0
4	{ $\Phi(t) \lambda(g^*t) p(t) p=c$ PIM}	-1175.5	6.33	0.03	0.04	15	-3038.0	-1206.5
5	{ $\Phi(t) \lambda(t) p(g^*t) p=c$ PIM}	-1174.8	7.07	0.02	0.03	16	-3039.4	-1208.0
6	{ $\Phi(t) \lambda(g^*t) p(g^*t) p=c$ PIM}	-1174.6	7.28	0.02	0.03	17	-3041.3	-1209.9
7	{ $\Phi(g^*t) \lambda(t) p(g^*t) p=c$ PIM}	-1173.2	8.59	0.01	0.01	17	-3040.0	-1208.6
8	{ $\Phi(g^*t) \lambda(g^*t) p(g^*t) p=c$ PIM}	-1172.6	9.28	0.01	0.01	18	-3041.5	-1210.1
9	{ $\Phi(t) \lambda(t) p(g) p=c$ }	-1172.0	9.83	0.00	0.01	10	-3023.9	-1192.5
10	{ $\Phi(.) \lambda(.) p(t) p=c$ PIM}	-1168.8	13.09	0.00	0.00	11	-3022.7	-1191.3
11	{ $\Phi(.) \lambda(t) p(t) p=c$ PIM}	-1158.4	23.41	0.00	0.00	11	-3012.4	-1181.0
12	{ $\Phi(.) \lambda(.) p(g) p=c$ }	-1156.9	24.96	0.00	0.00	9	-3006.7	-1175.3
13	{ $\Phi(g) \lambda(g) p(g) p=c$ }	-1156.5	25.29	0.00	0.00	11	-3010.5	-1179.1
14	{ $\phi(g^*t) \lambda(g^*t) p(g^*t) c(g^*t) F0(g^*t)$ PIM fail}	2.0	1183.85	0.00	0.00	1	0.0	0.0
15	{ $\phi(t) \lambda(t) p(t) c(t) F0(g^*t)$ PIM fail}	2.0	1183.85	0.00	0.00	1	0.0	0.0

Results Charron Lake: Capture rate



Results Charron Lake: Abundance and lambda



As with the Bog, the first survival rate estimate $\Phi(2015-17)$ had to be fixed at 1.0. In the following interval (2017-19) the survival rate fell dramatically: $\phi(2017-19) = 0.58 (0.14)$

Yet the Lambda stayed around 1.0 throughout (the estimate of overall Lambda from Model #2 is $\text{Lam}(\cdot) = 1.1 (0.14)$ with a 95% CI of (0.86, 1.41)) so recruitment must have been good in the latter interval to compensate for low survival. Pregnancy rates average 80% over the 3 years; being higher in 2017 (0.76 (2015), 0.87 (2017), 0.78 (2019)).

Conclusions

- Population sizes vary greatly among ranges, with Wabowden (Y1=105, Y2=147, Y3=121) and The Bog (Y1=144, Y2=211, Y3=78) being the smallest, Naosap-Reed (Y1=284, Y2=360, Y3=170) and Charron Lake (Y1=806, Y2=818, Y3=1160) being the largest.
- The ratio of M:F varied among populations from 0.6 to 0.9, with Charron Lake presenting a lower ratio (0.3-0.5).
- All populations except for Charron Lake, showed an increasing trend over the 2015-2017 period and a decreasing trend over the 2017-2019 period. Large drops in Lambda are not entirely accounted for by a decline in survival so there must have also been recruitment failures. Does this mean something global to all populations is happening, as opposed to specific environmental effects within individual populations.

